

Review of Probability

Random Variable

● Definition

Numerical characterization of outcome of a random event

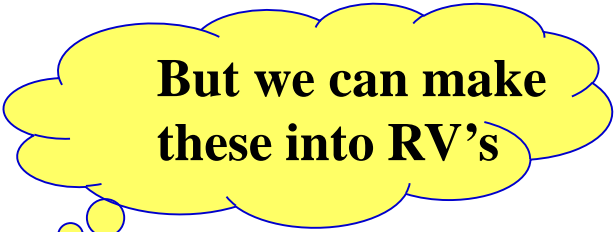
● Examples

- 1) Number on rolled dice
- 2) Temperature at specified time of day
- 3) Stock Market at close
- 4) Height of wheel going over a rocky road

Random Variable

- **Non-examples**

- 1) 'Heads' or 'Tails' on coin
- 2) Red or Black ball from urn



But we can make
these into RV's

- **Basic Idea** – don't know how to completely determine what value will occur
 - Can only specify probabilities of RV values occurring.

Two Types of Random Variables

Random Variable

```
graph TD; A[Random Variable] --> B[Discrete RV]; A --> C[Continuous RV];
```

Discrete RV

- Die
- Stocks

Continuous RV

- Temperature
- Wheel height

PDF for Continuous RV

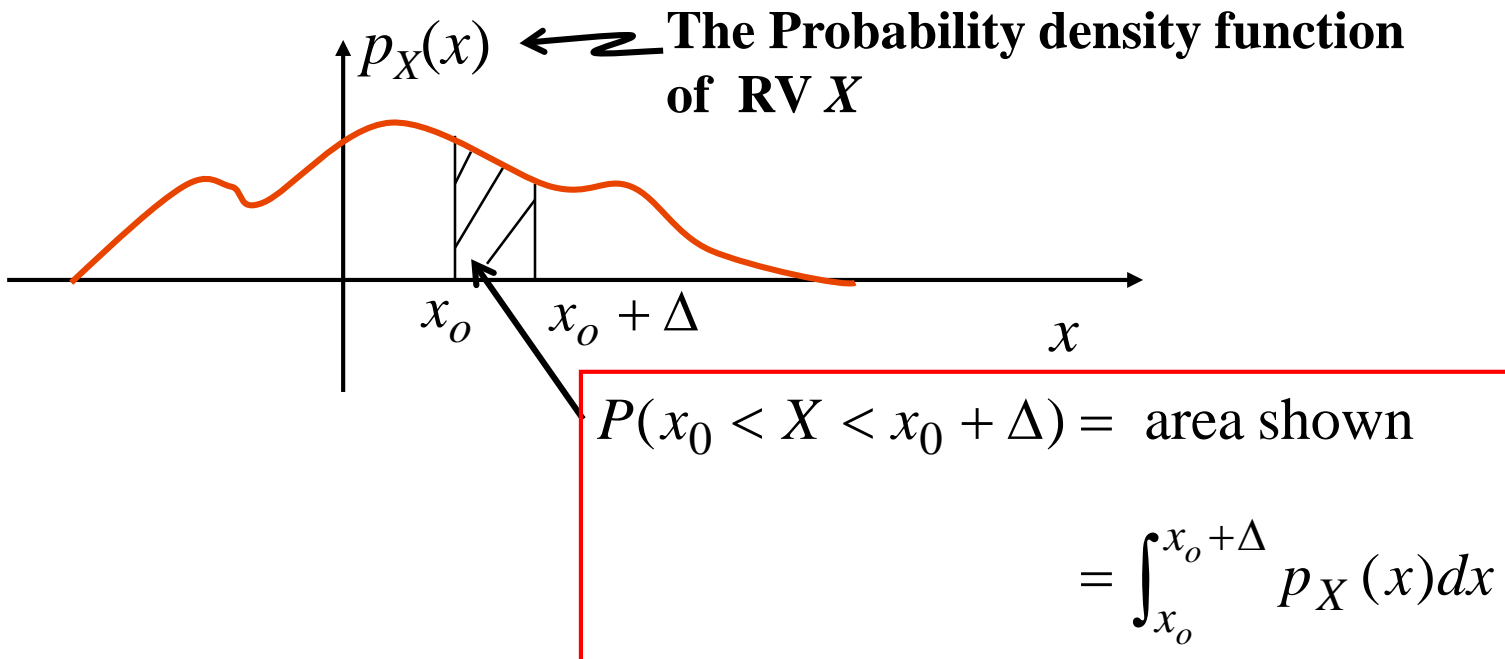
Given Continuous RV X ...

What is the probability that $X = x_0$?

➔ Oddity : $P(X = x_0) = 0$

Otherwise the Prob. “Sums” to infinity

➔ Need to think of Prob. *Density* Function (PDF)



Most Commonly Used PDF: Gaussian

A RV X with the following PDF is called a Gaussian RV

$$p_X(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-(x-m)^2 / 2\sigma^2}$$

m & σ are parameters of the Gaussian pdf

m = Mean of RV X

σ = Standard Deviation of RV X (Note: $\sigma > 0$)

σ^2 = Variance of RV X

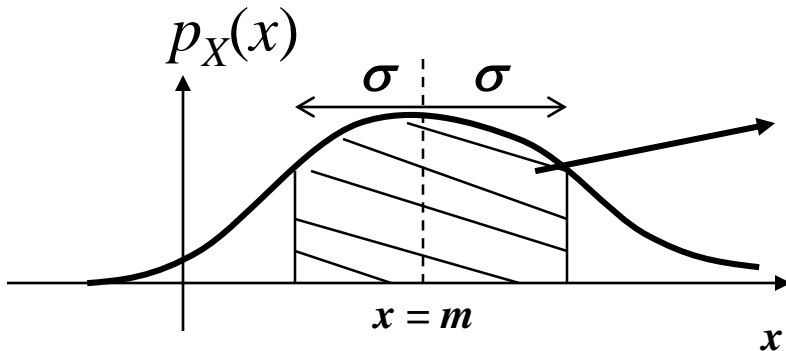
Notation: When X has Gaussian PDF we say $X \sim N(m, \sigma^2)$

Zero-Mean Gaussian PDF

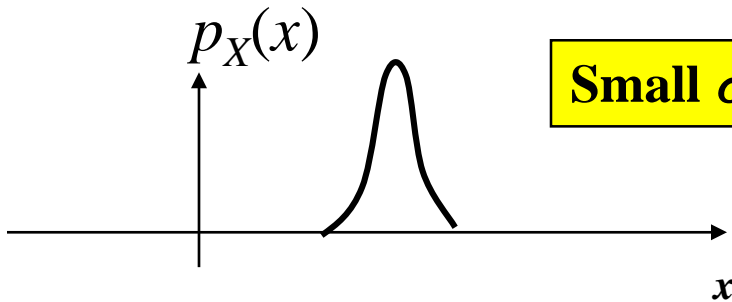
- Generally: take the noise to be Zero Mean

$$p_x(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-x^2/2\sigma^2}$$

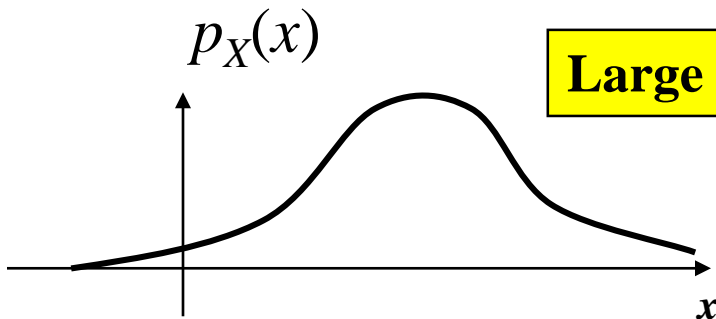
Effect of Variance on Gaussian PDF



Area within $\pm 1 \sigma$ of mean = 0.683
= 68.3%



Small Variability
(Small Uncertainty)



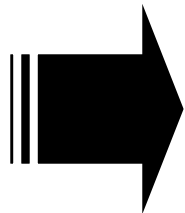
Large Variability
(Large Uncertainty)

Why Is Gaussian Used?

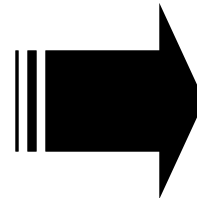
- Central Limit theorem (CLT)

The sum of N independent RVs has a pdf that tends to be Gaussian as $N \rightarrow \infty$

- So What! Here is what : Electronic systems generate internal noise due to random motion of electrons in electronic components. The noise is the result of summing the random effects of lots of electrons.



CLT applies



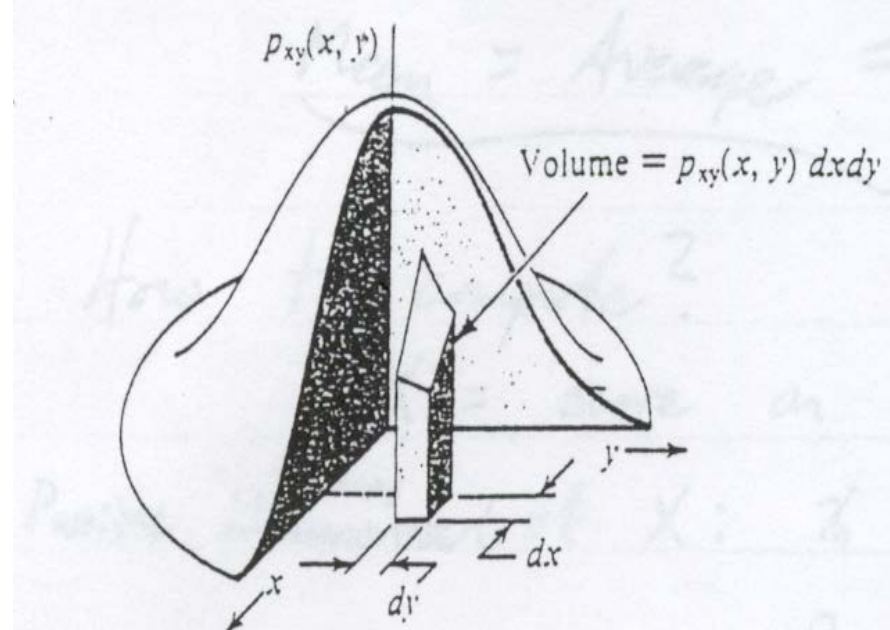
Guassian Noise

Joint PDF of RVs X and Y

$$p_{XY}(x, y)$$

Describes probabilities of joint events concerning X and Y . For example, the probability that X lies in interval $[a, b]$ and Y lies in interval $[a, b]$ is given by:

$$\Pr\{(a < X < b) \text{ and } (c < Y < d)\} = \int_a^b \int_c^d p_{XY}(x, y) dx dy$$



This graph shows a Joint PDF

Conditional PDF of Two RVs

When you have two RVs... often ask: What is the PDF of Y if X is constrained to take on a specific value.

In other words: What is the PDF of Y conditioned on the fact X is constrained to take on a specific value.

Ex.: Husband's salary X conditioned on wife's salary = \$100K?

First find all wives who make EXACTLY \$100K... how are their husband's salaries distributed.

Depends on the joint PDF because there are two RVs... but it should only depend on the slice of the joint PDF at $Y=\$100K$.

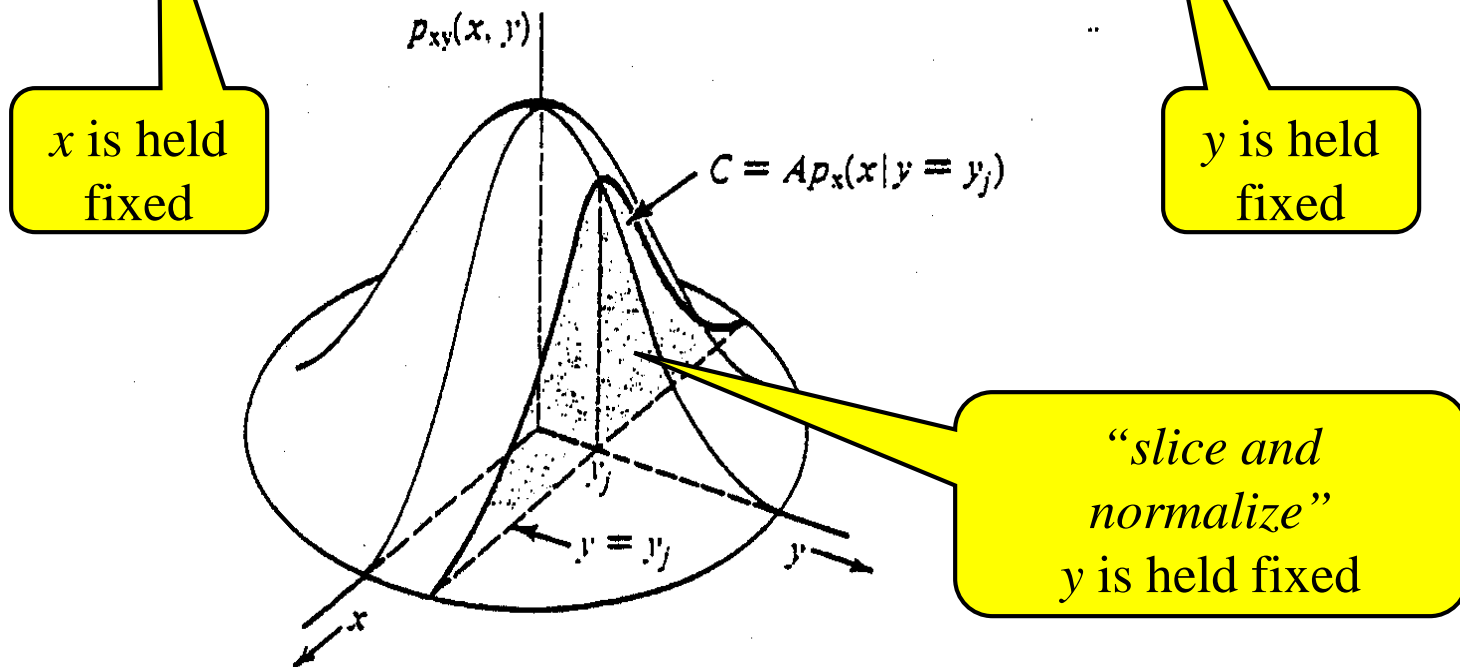
Now... we have to adjust this to account for the fact that the joint PDF (even its slice) reflects how likely it is that $Y=\$100K$ will occur (e.g., if $Y=10^5$ is unlikely then $p_{XY}(x, 10^5)$ will be small); so... if we divide by $p_Y(10^5)$ we adjust for this.

Conditional PDF (cont.)

Thus, the conditional PDFs are defined as (“slice and normalize”):

$$p_{Y|X}(y|x) = \begin{cases} \frac{p_{XY}(x,y)}{p_X(x)}, & p_X(x) \neq 0 \\ 0, & \text{otherwise} \end{cases}$$

$$p_{X|Y}(x|y) = \begin{cases} \frac{p_{XY}(x,y)}{p_Y(y)}, & p_Y(y) \neq 0 \\ 0, & \text{otherwise} \end{cases}$$



This graph shows a Conditional PDF

Independent RV's

Independence should be thought of as saying that:

Neither RV impacts the other statistically – thus, the values that one will likely take should be irrelevant to the value that the other has taken.

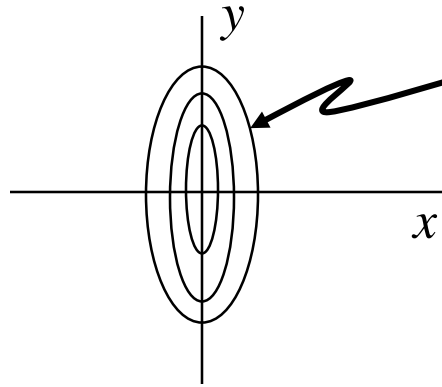
In other words: conditioning doesn't change the PDF!!!

$$p_{Y|X=x}(y | x) = \frac{p_{XY}(x, y)}{p_X(x)} = p_Y(y)$$

$$p_{X|Y=y}(x | y) = \frac{p_{XY}(x, y)}{p_Y(y)} = p_X(x)$$

Independent and Dependent Gaussian PDFs

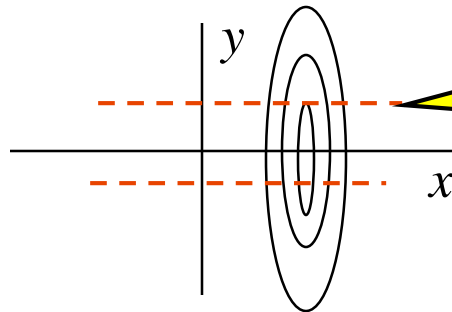
Independent
(zero mean)



Contours of $p_{XY}(x,y)$.

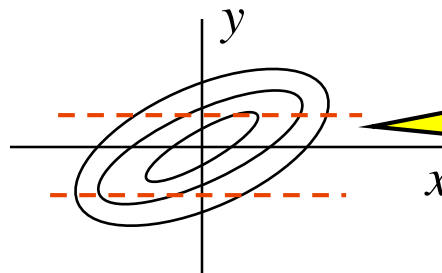
If X & Y are independent, then the contour ellipses are aligned with either the x or y axis

Independent
(non-zero mean)



Different slices give same normalized curves

Dependent



Different slices give different normalized curves

An “Independent RV” Result

RV's X & Y are independent if:

$$p_{XY}(x, y) = p_X(x)p_Y(y)$$

Here's why:

$$p_{Y|X=x}(y|x) = \frac{p_{XY}(x, y)}{p_X(x)} = \frac{\cancel{p_X(x)}p_Y(y)}{\cancel{p_X(x)}} = p_Y(y)$$

Characterizing RVs

- PDF tells everything about an RV
 - but sometimes they are “more than we need/know”
- So... we make due with a few Characteristics
 - Mean of an RV (Describes the centroid of PDF)
 - Variance of an RV (Describes the spread of PDF)
 - Correlation of RVs (Describes “tilt” of joint PDF)

Mean = Average = Expected Value

Symbolically: $E\{X\}$

Motivating Idea of Mean of RV

Motivation First w/ “Data Analysis View”

Consider RV $X =$ Score on a test Data: x_1, x_2, \dots, x_N

Possible values of RV X : $V_0 \ V_1 \ V_2 \dots \ V_{100}$
 $0 \ 1 \ 2 \ \dots \ 100$

$$\text{Test Average} = \bar{x} = \frac{\sum_{i=1}^N x_i}{N} = \frac{N_0 V_0 + N_1 V_1 + \dots + N_n V_{100}}{N} = \sum_{i=0}^{100} V_i \frac{N_i}{N}$$

$$N_i = \# \text{ of scores of value } V_i$$
$$N = \sum_{i=1}^n N_i \quad (\text{Total \# of scores})$$

$$\approx P(X = V_i)$$

This is called Data Analysis View

Statistics

But it motivates the Data Modeling View

Probability

Theoretical View of Mean

Data Analysis View leads to Probability Theory:

- For Discrete random Variables :

Data Modeling

$$E\{X\} = \sum_{n=1}^n x_i P_X(x_i)$$

Probability Function

- This Motivates form for Continuous RV:

$$E\{X\} = \int_{-\infty}^{\infty} x p_X(x) dx$$

Probability Density Function

Notation: $E\{X\} = \bar{X}$

Shorthand Notation

Aside: Probability vs. Statistics

Probability Theory

- » Given a PDF Model
- » Describe how the data will likely behave

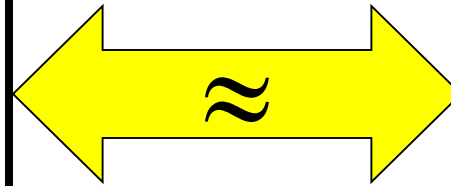
Statistics

- » Given a set of Data
- » Determine how the data did behave

$$E\{X\} = \int_{-\infty}^{\infty} x p_X(x) dx$$

PDF

Dummy Variable



“Law of Large Numbers”

$$Avg = \frac{1}{N} \sum_{i=1}^N x_i$$

Data

There is no DATA here!!!

The PDF models how the data will likely behave

There is no PDF here!!!

The Statistic measures how the data did behave

Variance of RV

There are similar Data vs. Theory Views here...

But let's go right to the theory!!

Variance: Characterizes how much you expect the RV to Deviate Around the Mean

$$\begin{aligned}\text{Variance: } \sigma^2 &= E\{(X - m_x)^2\} \\ &= \int (x - m_x)^2 p_X(x) dx\end{aligned}$$

Note : If zero mean...

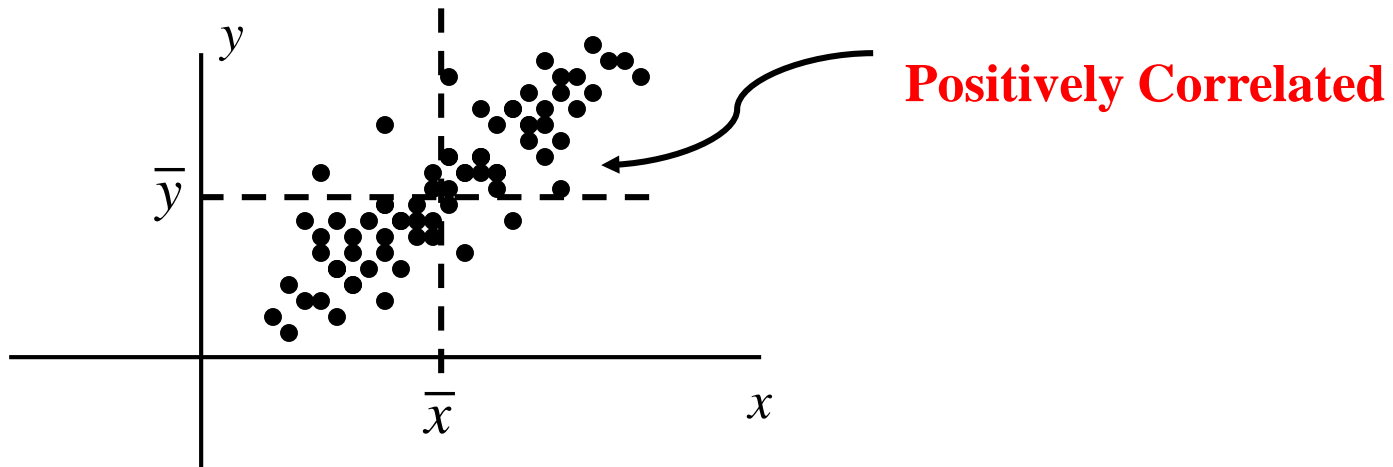
$$\begin{aligned}\sigma^2 &= E\{X^2\} \\ &= \int x^2 p_X(x) dx\end{aligned}$$

Motivating Idea of Correlation

Motivate First w/ Data Analysis View

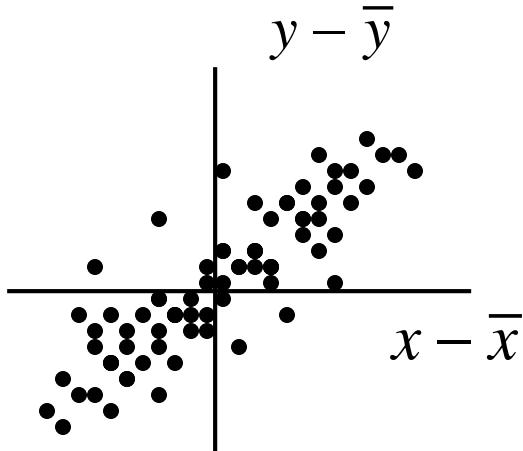
Consider a random experiment that observes the outcomes of two RVs:

Example: 2 RVs X and Y representing height and weight, respectively



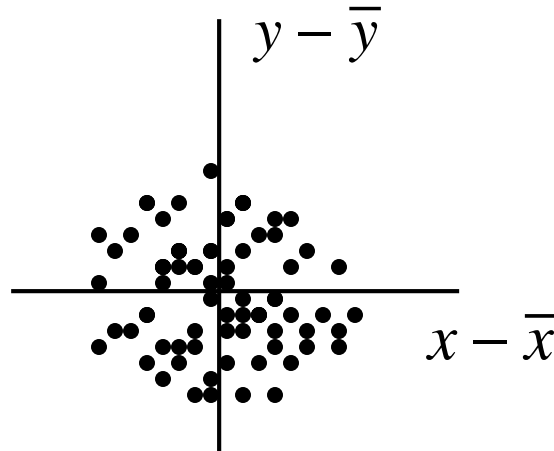
Illustrating 3 Main Types of Correlation

Data Analysis View:
$$C_{xy} = \frac{1}{N} \sum_{i=1}^N (x_i - \bar{x})(y_i - \bar{y})$$



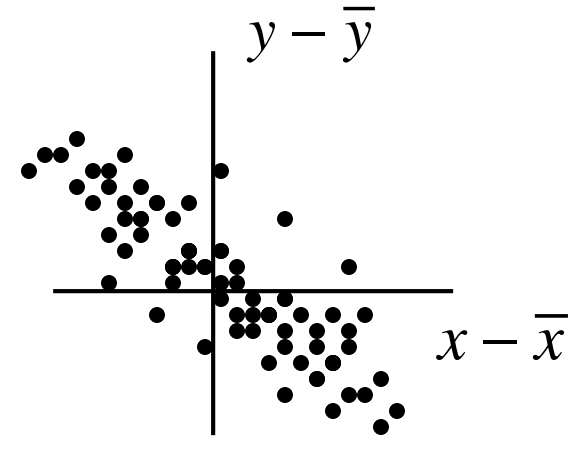
Positive Correlation
“Best Friends”

**GPA
&
Starting Salary**



Zero Correlation
i.e. uncorrelated
“Complete Strangers”

**Height
&
\$ in Pocket**



Negative Correlation
“Worst Enemies”

**Student Loans
&
Parents' Salary**

Prob. Theory View of Correlation

To capture this, define Covariance :

$$\sigma_{XY} = E\{(X - \bar{X})(Y - \bar{Y})\}$$

$$\sigma_{XY} = \iint (x - \bar{X})(y - \bar{Y}) p_{XY}(x, y) dx dy$$

If the RVs are both Zero-mean :

$$\sigma_{XY} = E\{XY\}$$

If $X = Y$:

$$\sigma_{XY} = \sigma_X^2 = \sigma_Y^2$$

If X & Y are independent, then:

$$\sigma_{XY} = 0$$

$$\text{If } \sigma_{XY} = E\{(X - \bar{X})(Y - \bar{Y})\} = 0$$

Then... Say that X and Y are **“uncorrelated”**

$$\text{If } \sigma_{XY} = E\{(X - \bar{X})(Y - \bar{Y})\} = 0$$

$$\text{Then } \underbrace{E\{XY\}} = \bar{X}\bar{Y}$$

Called **“Correlation of X & Y ”**

So... RVs X and Y are said to be uncorrelated

if $\sigma_{XY} = 0$

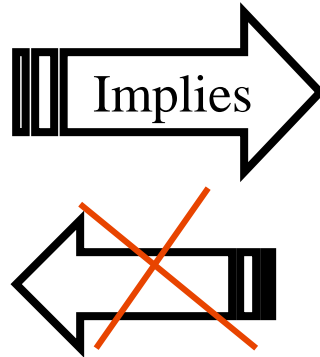
or equivalently... if $E\{XY\} = E\{X\}E\{Y\}$

Independence vs. Uncorrelated

X & Y are Independent

$$f_{XY}(x, y) = f_X(x)f_Y(y)$$

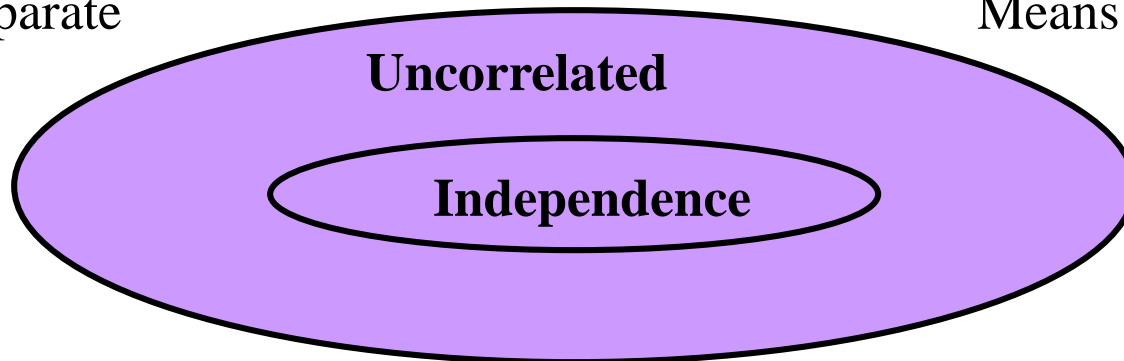
PDFs Separate



X & Y are Uncorrelated

$$E\{XY\} = E\{X\}E\{Y\}$$

Means Separate



INDEPENDENCE IS A STRONGER CONDITION !!!!

Confusing Covariance and Correlation Terminology

Covariance :

$$\sigma_{XY} = E\{(X - \bar{X})(Y - \bar{Y})\}$$

Correlation :

$$E\{XY\}$$

Same if zero mean



Correlation Coefficient :

$$\rho_{XY} = \frac{\sigma_{XY}}{\sigma_X \sigma_Y}$$

$$-1 \leq \rho_{XY} \leq 1$$

Covariance and Correlation For Random Vectors...

$$\mathbf{x} = [X_1 \ X_1 \ \cdots \ X_N]^T$$

Correlation Matrix :

$$\mathbf{R}_x = E\{\mathbf{x}\mathbf{x}^T\} = \begin{bmatrix} E\{X_1X_1\} & E\{X_1X_2\} & \cdots & E\{X_1X_N\} \\ E\{X_2X_1\} & E\{X_2X_2\} & \cdots & E\{X_2X_N\} \\ \vdots & \vdots & \ddots & \vdots \\ E\{X_NX_1\} & E\{X_NX_2\} & \cdots & E\{X_NX_N\} \end{bmatrix}$$

Covariance Matrix :

$$\mathbf{C}_x = E\{(\mathbf{x} - \bar{\mathbf{x}})(\mathbf{x} - \bar{\mathbf{x}})^T\}$$

A Few Properties of Expected Value

$$E\{X + Y\} = E\{X\} + E\{Y\}$$

$$E\{aX\} = aE\{X\}$$

$$E\{f(X)\} = \int f(x)p_X(x)dx$$

$$\text{var}\{X + Y\} = \begin{cases} \sigma_X^2 + \sigma_Y^2 + 2\sigma_{XY} \\ \sigma_X^2 + \sigma_Y^2, & \text{if } X \text{ \& } Y \text{ are uncorrelated} \end{cases}$$

$$\text{var}\{aX\} = a^2\sigma_X^2$$

$$\text{var}\{a + X\} = \sigma_X^2$$

$$\text{var}\{X + Y\} = E\{(X + Y - \bar{X} - \bar{Y})^2\}$$

$$= E\{(X_z + Y_z)^2\} \quad \text{where } X_z = X - \bar{X}$$

$$= E\{(X_z)^2 + (Y_z)^2 + 2X_zY_z\}$$

$$= E\{(X_z)^2\} + E\{(Y_z)^2\} + 2E\{X_zY_z\}$$

$$= \sigma_X^2 + \sigma_Y^2 + 2\sigma_{XY}$$

Joint PDF for Gaussian

Let $\mathbf{x} = [X_1 \ X_2 \ \dots \ X_N]^T$ be a vector of random variables. These random variables are said to be jointly Gaussian if they have the following PDF

$$p(\mathbf{x}) = \frac{1}{(2\pi)^{\frac{N}{2}} \sqrt{\det(\mathbf{C}_x)}} \exp\left\{-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_x)^T \mathbf{C}_x^{-1}(\mathbf{x} - \boldsymbol{\mu}_x)\right\}$$

where $\boldsymbol{\mu}_x$ is the mean vector and \mathbf{C}_x is the covariance matrix:

$$\boldsymbol{\mu}_x = E\{\mathbf{x}\} \quad \mathbf{C}_x = E\{(\mathbf{x} - \boldsymbol{\mu}_x)(\mathbf{x} - \boldsymbol{\mu}_x)^T\}$$

For the case of two jointly Gaussian RVs X_1 and X_2 with

$$E\{X_i\} = \mu_i \quad \text{var}\{X_i\} = \sigma_i^2 \quad E\{(X_1 - \mu_1)(X_2 - \mu_2)\} = \sigma_{12} \quad \rho = \sigma_{12}/(\sigma_1 \sigma_2)$$

Then...

$$p(x_1, x_2) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left\{-\frac{1}{2(1-\rho^2)}\left[\frac{(x_1 - \mu_1)^2}{\sigma_1^2} - 2\rho\frac{(x_1 - \mu_1)(x_2 - \mu_2)}{\sigma_1\sigma_2} + \frac{(x_2 - \mu_2)^2}{\sigma_2^2}\right]\right\}$$

It is easy to verify that X_1 and X_2 are uncorrelated (and independent!) if $\rho = 0$

Linear Transform of Jointly Gaussian RVs

Let $\mathbf{x} = [X_1 \ X_2 \ \dots \ X_N]^T$ be a vector of jointly Gaussian random variables with mean vector $\boldsymbol{\mu}_x$ and covariance matrix \mathbf{C}_x ...

Then the linear transform $\mathbf{y} = \mathbf{A}\mathbf{x} + \mathbf{b}$ is also jointly Gaussian with

$$\boldsymbol{\mu}_y = E\{\mathbf{y}\} = \mathbf{A}\boldsymbol{\mu}_x + \mathbf{b}$$

$$\mathbf{C}_y = E\{(\mathbf{y} - \boldsymbol{\mu}_y)(\mathbf{y} - \boldsymbol{\mu}_y)^T\} = \mathbf{A}\mathbf{C}_x\mathbf{A}^T$$

A special case of this is the sum of jointly Gaussian RVs... which can be handled using $\mathbf{A} = [1 \ 1 \ 1 \ \dots \ 1]$

Moments of Gaussian RVs

Let X be zero mean Gaussian with variance σ^2

Then the moments $E\{X^k\}$ are as follows:

$$E\{X^k\} = \begin{cases} 1 \cdot 3 \cdots (k-1) \sigma^k, & k \text{ even} \\ 0, & k \text{ odd} \end{cases}$$

Let $X_1 X_2 X_3 X_4$ be any four jointly Gaussian random variables with zero mean

Then...

$$E\{X_1 X_2 X_3 X_4\} = E\{X_1 X_2\} E\{X_3 X_4\} + E\{X_1 X_3\} E\{X_2 X_4\} + E\{X_1 X_4\} E\{X_2 X_3\}$$

Note that this can be applied to find $E\{X^2 Y^2\}$ if X and Y are jointly Gaussian

Chi-Squared Distribution

Let $X_1 X_2 \dots X_N$ be a set of zero-mean independent jointly Gaussian random variables each with unit variance.

Then the RV $Y = X_1^2 + X_2^2 + \dots + X_N^2$ is called a chi-squared (χ^2) RV of N degrees of freedom and has PDF given by

$$p(y) = \begin{cases} \frac{1}{2^{N/2} \Gamma(N/2)} y^{(N/2)-1} e^{-y/2}, & y \geq 0 \\ 0, & y < 0 \end{cases}$$

For this RV we have that:

$$E\{Y\} = N \quad \text{and} \quad \text{var}\{Y\} = 2N$$